# MAS3301 Bayesian Statistics: Project 

Semester 2

## 2008-9

## 1 Background

When analysing proportions (eg. with a binomial model) we can often use a beta prior distribution. However sometimes, particularly in more complicated cases, it is convenient to use a different form of prior distribution. Typically we transform the proportion to put it onto a $(-\infty, \infty)$ scale, rather than $(0,1)$, and then give the transformed proportion a normal distribution. In particular this makes it easy to give a relationship to two or more proportions in our prior distribution.

One form of transformation which is often used is known as "logits." If our proportion is $\theta$ then we transform this to

$$
\begin{equation*}
\eta=\log \left(\frac{\theta}{1-\theta}\right) \tag{1}
\end{equation*}
$$

and then give $\eta$ a normal prior distribution.
This project is concerned with analyses of this type.

## 2 Data

Each student will use a different data set. Each data set is identified by a reference number. A list of students and reference numbers and a separate list giving the data set for each reference number are provided in sections 5 and 6 below.

## 3 Tasks

1. Find an expression for the pdf of $\theta$ when $\eta$ is given by (1) and $\eta \sim N(m, v)$. You may wish to refer to Section 5.4 of the lecture notes.
2. Many people regard a beta $(1,1)$ distribution as a "noninformative" prior since the pdf is a constant. It might be thought that, if we use a logit transformation and then give $\eta$ a $N(0, v)$ prior distribution, then this will be "noninformative" if we make the variance $v$ large. However things are not as simple as this because, for large enough $v$, the prior distribution for $\theta$ becomes bimodal.
(a) Use R to plot graphs of the pdf of $\theta$ when $\eta$ is given by (1) and $\eta \sim N(0, v)$ for a range of values of $v$ in $1 \leq v \leq 4$ and hence deduce, at least approximately, the value of $v$ at which the distribution of $\theta$ becomes bimodal.
(b) Find analytically the exact value of $v$ at which the distribution of $\theta$ becomes bimodal when $\eta \sim N(0, v)$. Hint: The density of $\theta$ is symmetric about $\theta=1 / 2$ so look at the second derivative of the log density at this point.
3. Each patient in a sample of $n_{1}$ patients with a certain chronic illness is given a treatment. The number $x_{1}$ who show a particular response is recorded. Your values of $n_{1}$ and $x_{1}$ are in your data set. Given the value of a parameter $\theta_{1}$, we regard $x_{1}$ as an observation from the $\operatorname{binomial}\left(n_{1}, \theta_{1}\right)$ distribution. The value of $\theta_{1}$ is, however, unknown so we give it a prior
distribution by first transforming to $\eta_{1}$ using (1) and then giving $\eta_{1}$ a normal $n\left(m_{1}, v_{1}\right)$ prior distribution.
(a) In our prior beliefs, the lower quartile of $\theta_{1}$ is 0.4 and the upper quartile is 0.8 . Use these to find the lower and upper quartiles of $\eta_{1}$ and hence find the values of $m_{1}$ and $v_{1}$.
(b) Use numerical methods to find the posterior density of $\theta_{1}$ given your data and plot a graph showing both the prior and posterior densities.
(c) Find the posterior probability that $\theta_{1}>0.75$.
4. In another experiment with the same disease two further groups of patients are used. Group 2 contains $n_{2}$ patients and is given treatment $T_{2}$. Group 3 contains $n_{3}$ patients and is given treatment $T_{3}$. The number in Group $j$ who show the response is $x_{j}$. Given the values of parameters $\theta_{2}, \theta_{3}, x_{j}$ is regarded as an observation from the $\operatorname{binomial}\left(n_{j}, \theta_{j}\right)$ distribution and $x_{2}, x_{3}$ are independent.
Our prior distribution for $\theta_{2}, \theta_{3}$ is as follows. We transform $\theta_{2}, \theta_{3}$ to $\eta_{2}, \eta_{3}$ using (1) and then give $\eta_{1}, \eta_{2}$ a bivariate normal prior distribution with parameters as follows.

$$
\begin{aligned}
\mathrm{E}\left(\eta_{2}\right)=\mathrm{E}\left(\eta_{3}\right. & =0.4 \\
\operatorname{var}\left(\eta_{2}\right)=\operatorname{var}\left(\eta_{3}\right) & =0.6 \\
\operatorname{covar}\left(\eta_{2}, \eta_{3}\right) & =0.3
\end{aligned}
$$

(a) Use R to find and plot the joint posterior density of $\theta_{2}$ and $\theta_{3}$.
(b) Find and plot the prior and posterior probability density functions of the log relative risk $\gamma$ where

$$
\begin{aligned}
\gamma & =\log \left(\frac{\theta_{2} /\left(1-\theta_{2}\right)}{\theta_{3} /\left(1-\theta_{3}\right)}\right) \\
& =\log \left(\frac{\theta_{2}}{1-\theta_{2}}\right)-\log \left(\frac{\theta_{3}}{1-\theta_{3}}\right) \\
& =\eta_{2}-\eta_{3}
\end{aligned}
$$

(c) Comment on your results.

Hint: Part 4 of Practical 1 should help with Task 4. You will need to modify the $R$ functions appropriately.

## 4 Submission

Reports are to be submitted to the General Office of the School of Mathematics and Statistics no later than 4.00pm on Thursday 30th April (Week 10). As this counts for $10 \%$ of the module mark you will need to hand in your report at the reception window.

## 5 Data Set Reference Numbers

| Agnew, | Thomas | 1 |
| :---: | :---: | :---: |
| Armstrong, | Suzanne | 2 |
| Askew, | Leanne | 3 |
| Bannister, | John Jeffrey | 4 |
| Batey, | Aidan Joseph | 5 |
| Brooks, | Ciaran Anthony Foley | 6 |
| Brown, | Kate Elizabeth | 7 |
| Browne, | Colin John | 8 |
| Busby, | John | 9 |
| Chan, | Athena | 10 |
| Cheung, | Ho Yee | 11 |
| Cheyne-Vidal, | Nicola | 12 |
| Clarkson, | John-Frederic | 13 |
| Curwen, | Robert Anthony | 14 |
| D'Souza, | Francesca Kate | 15 |
| Dew, | Christopher | 16 |
| Egan, | Josephine | 17 |
| Ferguson, | Katy | 18 |
| Garvey, | Emma Jayne | 19 |
| Gill, | Martin | 20 |
| Goldthorpe, | Rowen | 21 |
| Gray, | Laura | 22 |
| Hallmark, | Laura Catherine | 23 |
| Jamison, | Deborah | 24 |
| Jayasuriya, | Gregory Anthony | 25 |
| Keene, | Stuart Peter | 26 |
| Kieselack, | Simon Nicholas | 27 |
| Lawson-Matthew, | Emma Jane | 28 |
| Longworth, | Jessica Daisy | 29 |
| MacGilchrist, | Graeme Alastair | 30 |
| Mann, | Kay Debby | 31 |
| McKinnon, | Alison | 32 |
| McParland, | Iain James | 33 |
| Millman, | Jill Fairless | 34 |
| Moncaster, | Sam | 35 |
| Munro, | Joseph | 36 |
| Nichols, | Ben | 37 |
| Payne, | Charlotte Elizabeth | 38 |
| Phillips, | Kate | 39 |
| Proom, | Rebecca Jane | 40 |
| Riley, | Anthony David | 41 |
| Roberts, | Catherine | 42 |
| Robertson, | Patrick | 43 |
| Smith, | James | 44 |
| Smith, | Warren Andrew | 45 |
| Tee, | Jane Katherine | 46 |
| Wang, | Yue | 47 |
| Wilkinson, | Nina | 48 |
| Wood, | Rachael Louise | 49 |
| Woodward, | Joe | 50 |
| Wrigley, | Amy | 51 |

## 6 Data Sets

| Ref.no. | $n_{1}$ | $x_{1}$ | $n_{2}$ | $x_{2}$ | $n_{3}$ | $x_{3}$ | Ref.no. | $n_{1}$ | $x_{1}$ | $n_{2}$ | $x_{2}$ | $n_{3}$ | $x_{3}$ |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | :--- | :--- | :--- | :--- | :--- |
| 1 | 25 | 18 | 55 | 32 | 20 | 9 | 34 | 27 | 22 | 53 | 28 | 48 | 41 |
| 2 | 26 | 19 | 54 | 28 | 21 | 16 | 35 | 28 | 20 | 52 | 32 | 47 | 37 |
| 3 | 27 | 24 | 53 | 30 | 22 | 17 | 36 | 29 | 21 | 51 | 33 | 46 | 36 |
| 4 | 28 | 24 | 52 | 28 | 23 | 16 | 37 | 30 | 24 | 50 | 25 | 45 | 36 |
| 5 | 29 | 25 | 51 | 33 | 24 | 19 | 38 | 31 | 19 | 49 | 33 | 44 | 36 |
| 6 | 30 | 20 | 50 | 30 | 25 | 21 | 39 | 32 | 24 | 48 | 24 | 43 | 29 |
| 7 | 31 | 25 | 49 | 36 | 26 | 20 | 40 | 33 | 23 | 47 | 24 | 42 | 28 |
| 8 | 32 | 23 | 48 | 32 | 27 | 23 | 41 | 34 | 28 | 46 | 19 | 41 | 31 |
| 9 | 33 | 21 | 47 | 27 | 28 | 23 | 42 | 35 | 28 | 45 | 29 | 40 | 34 |
| 10 | 34 | 28 | 46 | 19 | 29 | 21 | 43 | 36 | 28 | 44 | 30 | 39 | 29 |
| 11 | 35 | 22 | 45 | 35 | 30 | 26 | 44 | 37 | 27 | 43 | 23 | 38 | 31 |
| 12 | 36 | 25 | 44 | 19 | 31 | 24 | 45 | 38 | 30 | 42 | 23 | 37 | 31 |
| 13 | 37 | 24 | 43 | 24 | 32 | 29 | 46 | 39 | 32 | 41 | 25 | 36 | 27 |
| 14 | 38 | 30 | 42 | 24 | 33 | 29 | 47 | 40 | 28 | 40 | 24 | 35 | 28 |
| 15 | 39 | 30 | 41 | 27 | 34 | 27 | 48 | 41 | 31 | 39 | 25 | 34 | 24 |
| 16 | 40 | 30 | 40 | 22 | 35 | 27 | 49 | 42 | 34 | 38 | 23 | 33 | 27 |
| 17 | 41 | 28 | 39 | 19 | 36 | 31 | 50 | 43 | 35 | 37 | 23 | 32 | 27 |
| 18 | 42 | 34 | 38 | 19 | 37 | 28 | 51 | 44 | 30 | 36 | 24 | 31 | 24 |
| 19 | 43 | 31 | 37 | 23 | 38 | 30 | 52 | 45 | 29 | 35 | 23 | 30 | 28 |
| 20 | 44 | 31 | 36 | 21 | 39 | 29 | 53 | 46 | 35 | 34 | 19 | 29 | 25 |
| 21 | 45 | 39 | 35 | 18 | 40 | 28 | 54 | 47 | 34 | 33 | 17 | 28 | 27 |
| 22 | 46 | 34 | 34 | 19 | 41 | 34 | 55 | 48 | 40 | 32 | 15 | 27 | 22 |
| 23 | 47 | 33 | 33 | 23 | 42 | 30 | 56 | 49 | 34 | 31 | 23 | 26 | 16 |
| 24 | 48 | 35 | 32 | 24 | 43 | 31 | 57 | 50 | 40 | 30 | 19 | 25 | 22 |
| 25 | 49 | 36 | 31 | 15 | 44 | 34 | 58 | 51 | 33 | 29 | 16 | 24 | 18 |
| 26 | 50 | 44 | 30 | 22 | 45 | 37 | 59 | 52 | 37 | 28 | 16 | 23 | 16 |
| 27 | 51 | 34 | 29 | 21 | 46 | 37 | 60 | 53 | 38 | 27 | 16 | 22 | 17 |
| 28 | 52 | 42 | 28 | 18 | 47 | 36 | 61 | 54 | 39 | 26 | 14 | 21 | 20 |
| 29 | 53 | 44 | 27 | 15 | 48 | 39 | 62 | 55 | 38 | 25 | 19 | 20 | 18 |
| 30 | 54 | 40 | 26 | 18 | 49 | 34 | 63 | 31 | 27 | 34 | 15 | 33 | 28 |
| 31 | 55 | 41 | 25 | 14 | 50 | 42 | 64 | 32 | 25 | 33 | 17 | 34 | 27 |
| 32 | 25 | 21 | 55 | 35 | 50 | 41 | 65 | 33 | 26 | 32 | 19 | 33 | 26 |
| 33 | 26 | 23 | 54 | 34 | 49 | 42 | 66 | 34 | 26 | 31 | 16 | 34 | 28 |

