MAS2903 Problems Class 1

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Semester 2, 2019—2020

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- Assignment 1
 - Due in by 3pm, Friday 6th March
 - Consists of questions 4, 9, 10, 11, 12
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- Next week: computer practical on Tuesday at 12pm, to try out some of the R questions in Chapter 5

Question 5

An experiment can result in four possible outcomes x_1 , x_2 , x_3 and x_4 .

The probability of each of these outcomes is affected by a parameter θ which can take one of three values: θ_1 , θ_2 and θ_3 :

$p(x \theta)$	<i>X</i> ₁	<i>X</i> ₂	<i>X</i> ₃	<i>X</i> ₄
θ_1	0.1	0.2	0.3	0.4
θ_2	0.3	0.3	0.2	0.2
θ_3	0.5	0.1	0.1	0.3

Question 5

The prior plausibilities attaching to θ_1 , θ_2 and θ_3 are 0.1, 0.4 and 0.5.

Find the posterior probability function for θ after observing x_1 . Repeat the calculations for x_2 , x_3 and x_4 .

$$Pr(\theta_1|x_1) = \frac{Pr(x_1|\theta_1)Pr(\theta_1)}{Pr(x_1|\theta_1)Pr(\theta_1) + Pr(x_1|\theta_2)Pr(\theta_2) + Pr(x_1|\theta_3)Pr(\theta_3)}$$

$$Pr(\theta_1|x_1) = \frac{Pr(x_1|\theta_1)Pr(\theta_1)}{Pr(x_1|\theta_1)Pr(\theta_1) + Pr(x_1|\theta_2)Pr(\theta_2) + Pr(x_1|\theta_3)Pr(\theta_3)}$$
$$= \frac{0.1 \times 0.1}{0.1 \times 0.1 + 0.3 \times 0.4 + 0.5 \times 0.5}$$

$$Pr(\theta_{1}|x_{1}) = \frac{Pr(x_{1}|\theta_{1})Pr(\theta_{1})}{Pr(x_{1}|\theta_{1})Pr(\theta_{1}) + Pr(x_{1}|\theta_{2})Pr(\theta_{2}) + Pr(x_{1}|\theta_{3})Pr(\theta_{3})}$$

$$= \frac{0.1 \times 0.1}{0.1 \times 0.1 + 0.3 \times 0.4 + 0.5 \times 0.5}$$

$$= 0.02632,$$

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$$= \frac{0.1 \times 0.1}{0.1 \times 0.1 + 0.3 \times 0.4 + 0.5 \times 0.5}$$

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$$Pr(\theta_2|x_1) = \frac{Pr(x_1|\theta_2)Pr(\theta_2)}{Pr(x_1|\theta_1)Pr(\theta_1) + Pr(x_1|\theta_2)Pr(\theta_2) + Pr(x_1|\theta_3)Pr(\theta_3)}$$

$$Pr(\theta_1|x_1) = \frac{Pr(x_1|\theta_1)Pr(\theta_1)}{Pr(x_1|\theta_1)Pr(\theta_1) + Pr(x_1|\theta_2)Pr(\theta_2) + Pr(x_1|\theta_3)Pr(\theta_3)}$$

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$$= \frac{0.3 \times 0.4}{0.1 \times 0.1 + 0.3 \times 0.4 + 0.5 \times 0.5}$$

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$$= \frac{0.1 \times 0.1}{0.1 \times 0.1 + 0.3 \times 0.4 + 0.5 \times 0.5}$$

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$$= \frac{0.3 \times 0.4}{0.1 \times 0.1 + 0.3 \times 0.4 + 0.5 \times 0.5}$$

$$= 0.31579,$$

and

$$Pr(\theta_3|x_1) = 1 - Pr(\theta_1|x_1) - Pr(\theta_2|x_1)$$

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= 1 - 0.02632 - 0.31579
= 0.65789.

Thus, after observing x_1 , we have the following probability function for θ :

and

$$Pr(\theta_3|x_1) = 1 - Pr(\theta_1|x_1) - Pr(\theta_2|x_1)$$

= 1 - 0.02632 - 0.31579
= 0.65789.

Thus, after observing x_1 , we have the following probability function for θ :

θ	θ_1	θ_2	θ_3
$Pr(\theta = \theta_i x_1)$	0.02632	0.31579	0.65789

Same for x_2 , x_3 and x_4 .

Question 6

Extensive manuscripts of two scribes show that they differ on the relative frequencies with which they use two alternative vowel forms e and oe, as in the following table:

	е	œ
Scribe A	0.3	0.7
Scribe B	0.5	0.5

Question 6

A new manuscript fragment, known to be by one of the scribes, uses the vowel form five times, three times as e and twice as e.

If an independent assessment of similar fragments favours Scribe A as author with probability 0.7, show that the evidence from the vowel form is sufficient to turn the odds in favour of Scribe B.

Let X = number of e's observed.

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Then

$$Pr(\theta_A)=0.7, \quad Pr(\theta_B)=0.3,$$
 $X|\theta_A\sim Bin(5,0.3) \quad {\rm and} \quad X|\theta_B\sim Bin(5,0.5).$

Hence

$$Pr(X = 3|\theta_A) = {5 \choose 3} 0.3^3 \times 0.7^2 = 0.1323$$

Hence

$$Pr(X = 3|\theta_A) = {5 \choose 3} 0.3^3 \times 0.7^2 = 0.1323$$

$$Pr(X = 3|\theta_B) = {5 \choose 3} 0.5^3 \times 0.5^2 = 0.3125.$$

$$Pr(\theta_A|x=3) = \frac{Pr(X=3|\theta_A)Pr(\theta_A)}{Pr(X=3|\theta_A)Pr(\theta_A) + Pr(X=3|\theta_B)Pr(\theta_B)}$$

$$Pr(\theta_A|x=3) = \frac{Pr(X=3|\theta_A)Pr(\theta_A)}{Pr(X=3|\theta_A)Pr(\theta_A) + Pr(X=3|\theta_B)Pr(\theta_B)}$$
$$= \frac{0.1323 \times 0.7}{0.1323 \times 0.7 + 0.3125 \times 0.3}$$

$$Pr(\theta_A|x=3) = \frac{Pr(X=3|\theta_A)Pr(\theta_A)}{Pr(X=3|\theta_A)Pr(\theta_A) + Pr(X=3|\theta_B)Pr(\theta_B)}$$

$$= \frac{0.1323 \times 0.7}{0.1323 \times 0.7 + 0.3125 \times 0.3}$$

$$= 0.497 (3 d.p.)$$

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$$= \frac{0.1323 \times 0.7}{0.1323 \times 0.7 + 0.3125 \times 0.3}$$

$$= 0.497 (3 d.p.)$$

$$Pr(\theta_B|x=3) = 1 - Pr(\theta_A|x=3) = 0.503 (3 d.p.).$$

Therefore, using Bayes' Theorem

$$Pr(\theta_A|x=3) = \frac{Pr(X=3|\theta_A)Pr(\theta_A)}{Pr(X=3|\theta_A)Pr(\theta_A) + Pr(X=3|\theta_B)Pr(\theta_B)}$$

$$= \frac{0.1323 \times 0.7}{0.1323 \times 0.7 + 0.3125 \times 0.3}$$

$$= 0.497 (3 d.p.)$$

$$Pr(\theta_B|x=3) = 1 - Pr(\theta_A|x=3) = 0.503 \ (3 d.p.).$$

Thus, the odds have turned from 7 : 3 in favour of Scribe A to 503 : 497 in favour of Scribe B.

Question 7

If X_1, X_2, \dots, X_n are independent random variables with probability function

$$f(x|\theta) = \theta(1-\theta)^{x-1}, \quad x = 1, 2, ...,$$

show that $T = \sum_{i=1}^{n} X_i$ is sufficient for θ .

$$f(\mathbf{x}|\theta) = \prod_{i=1}^{n} \theta (1-\theta)^{x_i-1}$$

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$$= \theta^n (1-\theta)^{x_1-1} (1-\theta)^{x_2-1} \times \ldots \times (1-\theta)^{x_n-1}$$

$$f(\mathbf{x}|\theta) = \prod_{i=1}^{n} \theta (1-\theta)^{x_{i}-1}$$

$$= \theta^{n} (1-\theta)^{x_{1}-1} (1-\theta)^{x_{2}-1} \times \dots \times (1-\theta)^{x_{n}-1}$$

$$= \theta^{n} (1-\theta)^{\sum x_{i}-n}$$

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$$= \theta^n (1-\theta)^{\sum x_i-n}$$

$$= 1 \times \theta^n (1-\theta)^{\sum x_i-n}$$

$$f(\mathbf{x}|\theta) = \prod_{i=1}^{n} \theta (1-\theta)^{x_{i}-1}$$

$$= \theta^{n} (1-\theta)^{x_{1}-1} (1-\theta)^{x_{2}-1} \times \dots \times (1-\theta)^{x_{n}-1}$$

$$= \theta^{n} (1-\theta)^{\sum x_{i}-n}$$

$$= 1 \times \theta^{n} (1-\theta)^{\sum x_{i}-n}$$

$$= h(\mathbf{x}) g\left(\sum x_{i}, \theta\right)$$

The joint probability function is

$$f(\mathbf{x}|\theta) = \prod_{i=1}^{n} \theta (1-\theta)^{x_{i}-1}$$

$$= \theta^{n} (1-\theta)^{x_{1}-1} (1-\theta)^{x_{2}-1} \times \dots \times (1-\theta)^{x_{n}-1}$$

$$= \theta^{n} (1-\theta)^{\sum x_{i}-n}$$

$$= 1 \times \theta^{n} (1-\theta)^{\sum x_{i}-n}$$

$$= h(\mathbf{x}) g\left(\sum x_{i}, \theta\right)$$

where $h(\mathbf{x}) = 1$ and $g(t, \theta) = \theta^n (1 - \theta)^{t-n}$. Therefore, by the Factorisation Theorem, $T = \sum X_i$ is sufficient for θ .

Question 8

If X_1, X_2, \dots, X_n are independent $N(5, \sigma^2)$ random variables, show that

$$T = \sum_{i=1}^{n} (X_i - 5)^2$$

is sufficient for σ .

$$f(\mathbf{x}|\theta) = \prod_{i=1}^{n} \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left\{-\frac{(x_i - 5)^2}{2\sigma^2}\right\}$$

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$$f(\mathbf{x}|\theta) = \prod_{i=1}^{n} \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left\{-\frac{(x_i - 5)^2}{2\sigma^2}\right\}$$

$$= (2\pi)^{-n/2} \sigma^{-n} \exp\left\{-\frac{1}{2\sigma^2} \sum_{i=1}^{n} (x_i - 5)^2\right\}$$

$$= 1 \times (2\pi)^{-n/2} \sigma^{-n} \exp\left\{-\frac{1}{2\sigma^2} \sum_{i=1}^{n} (x_i - 5)^2\right\}$$

$$f(\mathbf{x}|\theta) = \prod_{i=1}^{n} \frac{1}{\sqrt{2\pi\sigma^{2}}} \exp\left\{-\frac{(x_{i}-5)^{2}}{2\sigma^{2}}\right\}$$

$$= (2\pi)^{-n/2}\sigma^{-n} \exp\left\{-\frac{1}{2\sigma^{2}}\sum_{i=1}^{n}(x_{i}-5)^{2}\right\}$$

$$= 1 \times (2\pi)^{-n/2}\sigma^{-n} \exp\left\{-\frac{1}{2\sigma^{2}}\sum_{i=1}^{n}(x_{i}-5)^{2}\right\}$$

$$= h(\mathbf{x}) g\left(\sum (x_{i}-5)^{2}, \sigma\right)$$

The joint probability density function is

$$f(\mathbf{x}|\theta) = \prod_{i=1}^{n} \frac{1}{\sqrt{2\pi\sigma^{2}}} \exp\left\{-\frac{(x_{i}-5)^{2}}{2\sigma^{2}}\right\}$$

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$$= 1 \times (2\pi)^{-n/2}\sigma^{-n} \exp\left\{-\frac{1}{2\sigma^{2}}\sum_{i=1}^{n}(x_{i}-5)^{2}\right\}$$

$$= h(\mathbf{x})g\left(\sum(x_{i}-5)^{2},\sigma\right)$$
where $h(\mathbf{x}) = 1$ and $g(t,\sigma) = (2\pi)^{-n/2}\sigma^{-n}e^{-t/(2\sigma^{2})}$.

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The joint probability density function is

$$f(\mathbf{x}|\theta) = \prod_{i=1}^{n} \frac{1}{\sqrt{2\pi\sigma^{2}}} \exp\left\{-\frac{(x_{i}-5)^{2}}{2\sigma^{2}}\right\}$$

$$= (2\pi)^{-n/2} \sigma^{-n} \exp\left\{-\frac{1}{2\sigma^{2}} \sum_{i=1}^{n} (x_{i}-5)^{2}\right\}$$

$$= 1 \times (2\pi)^{-n/2} \sigma^{-n} \exp\left\{-\frac{1}{2\sigma^{2}} \sum_{i=1}^{n} (x_{i}-5)^{2}\right\}$$

$$= h(\mathbf{x}) g\left(\sum (x_{i}-5)^{2}, \sigma\right)$$

where
$$h(\mathbf{x}) = 1$$
 and $g(t, \sigma) = (2\pi)^{-n/2} \sigma^{-n} e^{-t/(2\sigma^2)}$.

Therefore, by the Factorisation Theorem, $T = \sum (X_i - 5)^2$ is sufficient for σ .